

# **STA 35C: Statistical Data Science III**

## **Lecture 19: Multiple Hypotheses Testing (cont'd) + Review for Midterm 2**

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# Announcement

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## **Midterm 2** on Fri, May 16 (12:10 pm–1:00 pm in class)

- See [Canvas announcement](#) (or Lec. 17/18 slides) for allowed materials, etc.
- Coverage: Lectures 12–19
- A [practice midterm](#) and [answer key](#) are available on the course webpage
- Office hours this week:
  - Instructor: Wed, 4–6pm (extended); no OH Thu
  - TA: Thu 1–2pm

## **Remote lecture (Zoom)** on Mon, May 19

- Zoom link will be emailed via Canvas

# Today's topics

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- **Recap: Multiple hypotheses testing**
  - Goals to control false positives
- **Brief review for midterm 2**
  - Cross-validation
  - Bootstrap
  - Subset selection
  - Regularization
  - Multiple hypotheses testing

## Recap: Multiple testing

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### Single-hypothesis test:

- Typically set up  $H_0$ , and gather data to reject it if there is significant evidence
- Type I error = false positive; Type II error = false negative
- Each test has Type I error at most  $\alpha$  (e.g. 0.05)

### Modern data analysis: multiple tests simultaneously

- E.g. Testing thousands of predictors or biomarkers to discovery significant ones
- If  $m$  is large, false rejections can occur easily by chance
- On average,  $\alpha \times m$  false positives if each is tested at level  $\alpha$

**Key challenge:** How to address inflated false positives as  $m$  grows

# Hypothesis testing as classification

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A single hypothesis test classifies  $H_0$  as “true or not”:

- **Goal:** Discover “real phenomenon” ( $H_1$ ) or conclude non-existence ( $H_0$ )
  - $H_0$  is true  $\iff$  no real effect
  - $H_0$  is false  $\iff$  there is a real effect ( $H_1$ )
  - We “discover” an effect by rejecting  $H_0$
- **Test as classification:** Depending on evidence gathered from data,
  - Reject  $H_0 \iff$  classify  $\hat{H} = 1$
  - Fail to reject  $H_0 \iff$  classify  $\hat{H} = 0$

	$H_0$ is true (" <b>H=0</b> ")	$H_0$ is not true (" <b>H=1</b> ")
Reject $H_0$ (" <b><math>\hat{H}=1</math></b> ")	FP (Type I)	TP
Not reject $H_0$ (" <b><math>\hat{H}=0</math></b> ")	TN	FN (Type II)

## Hypothesis test at level $\alpha$

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Consider the probabilities of each outcome for hypothesis test

	$H_0$ is true (" $H=0$ ")	$H_0$ is not true (" $H=1$ ")
Reject $H_0$ (" $\hat{H}=1$ ")	$p_{FP}$	$p_{TP}$
Not reject $H_0$ (" $\hat{H}=0$ ")	$p_{TN}$	$p_{FN}$

Hypothesis test at level  $\alpha$ :

- $\Pr(\text{reject } H_0 \mid H_0 \text{ true}) \leq \alpha$
- That is, the chance of a false positive is at most  $\alpha$

$$\Pr(\hat{H} = 1 \mid H = 0) = \frac{\Pr(\hat{H} = 1 \ \& \ H = 0)}{\Pr(H = 0)} = \frac{p_{FP}}{p_{FP} + p_{TN}} \leq \alpha$$

## Testing multiple hypotheses at level $\alpha$

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Suppose we test  $m$  hypotheses  $H_{0,1}, \dots, H_{0,m}$ , all at level  $\alpha$ , obtaining confusion matrix:

	$H_0$ is true	$H_0$ is not true
Reject $H_0$	$N_{\text{FP}}$	$N_{\text{TP}}$
Not reject $H_0$	$N_{\text{TN}}$	$N_{\text{FN}}$

- $N_{\text{FP}}, N_{\text{TP}}, N_{\text{TN}}, N_{\text{FN}}$  are random variables that sum to  $m$
- Roughly, we expect  $N_{\text{FP}} \approx m \cdot p_{\text{FP}}$ ; when all  $m$  nulls are true,  $N_{\text{FP}} \approx m \cdot \alpha$

If these  $m$  tests are independent,

- Probability of *at least one* false positive  $\approx 1 - (1 - \alpha)^m$
- For  $m = 20, \alpha = 0.05$ , that probability is  $\approx 64\%$

## Family-wise error rate (FWER)

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	$H_0$ is true	$H_0$ is not true
Reject $H_0$	$N_{\text{FP}}$	$N_{\text{TP}}$
Not reject $H_0$	$N_{\text{TN}}$	$N_{\text{FN}}$

**Goal:** Ensure  $N_{\text{FP}} < 1$  with high probability

$$\text{FWER} = \Pr(N_{\text{FP}} \geq 1)$$

- Bonferroni correction sets each test at  $\alpha/m$  to keep  $\text{FWER} \leq \alpha$  (union bound)
- Holm's step-down procedure refines this by adapting thresholds step by step

**Interpretation:** Controlling  $\text{FWER} \leq \alpha$  ensures we have *no* Type I errors with probability at least  $1 - \alpha$



## False discovery rate (FDR)

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	$H_0$ is true	$H_0$ is not true
Reject $H_0$	$N_{\text{FP}}$	$N_{\text{TP}}$
Not reject $H_0$	$N_{\text{TN}}$	$N_{\text{FN}}$

**FDR Strategy:** Increase  $N_{\text{TP}}$  at the cost of tolerating a moderate  $N_{\text{FP}}$

- Strict FWER control often yields many Type II errors (missing real signals)
- FDR-based approach lets us accept some false positives but aims for higher power (detecting more TP)
  - $N_{\text{FP}}$ : “false discoveries”
  - $N_{\text{TP}}$ : “true discoveries”

## False discovery rate control

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	$H_0$ is true	$H_0$ is not true
Reject $H_0$	$N_{\text{FP}}$	$N_{\text{TP}}$
Not reject $H_0$	$N_{\text{TN}}$	$N_{\text{FN}}$

**False discovery proportion:** fraction of false discoveries among all "claimed" ( $\hat{H} = 1$ )

$$\text{FDP} = \frac{N_{\text{FP}}}{N_{\text{FP}} + N_{\text{TP}}}$$

**False discovery rate (FDR):**  $\text{FDR} = \mathbb{E}[\text{FDP}]$

- Controlling FDR at  $q$  (e.g., 5% or 10%) means  $\mathbb{E}[\text{FDP}] \leq q$
- Methods like Benjamini–Hochberg aim to maintain  $\text{FDR} \leq q$  while rejecting more nulls than strict FWER approaches

## Pop-up quiz: Comparing FDR vs. FWER

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You have  $m$  hypotheses to test. The *False Discovery Rate* (FDR) is defined as  $\mathbb{E}[\text{FDP}]$ , where  $\text{FDP} = \frac{\#FP}{\#FP + \#TP}$ . **Which statement best captures differences between FDR and FWER?**

- (A) FDR forces the probability of zero false positives to stay below  $\alpha$ , whereas FWER allows a small fraction  $q$ .
- (B) FDR aims to keep  $\mathbb{E}[\text{fraction of false positives among rejections}] \leq q$ , while FWER demands  $\Pr(\text{at least one false positive}) \leq \alpha$ .
- (C) Under FDR control, no false positives are allowed once you discover enough true positives.
- (D) FDR only works for independent tests, but FWER can handle correlated tests without adjustments.

**Answer: (B).**

FDR control (e.g., Benjamini–Hochberg) allows a certain fraction of false positives *on average*, whereas FWER control (e.g., Bonferroni/Holm) requires the chance of *any* false positive be controlled below  $\alpha$ .

# Review: Cross-validation

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**Goal:** Estimate test performance from training data alone

**Key ideas:**

- Single split (validation set): random partition into train/test; simple but high variance
- LOOCV (leave-one-out): train on  $n - 1$  points, validate on 1 point, repeat for all points
- $k$ -fold CV: partition data into  $k$  folds, systematically rotate which fold is the validation set

**Trade-offs:**

- Fewer folds (e.g. 5- or 10-fold) reduce computation but can have slightly higher variance
- LOOCV uses maximum training size ( $n - 1$ ) but is more expensive and can have higher correlation across folds

**Usage:**

- Model selection: pick model that yields lowest CV error
- Tuning parameters (e.g.  $\lambda$  in ridge/lasso)

# Review: Bootstrap

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**Goal:** Approximate the sampling distribution (e.g. standard errors) using just one dataset

**Method:**

- Sample  $n$  points *with replacement* from the original dataset of size  $n$  (a “bootstrap sample”)
- Compute desired statistic (mean, regression coefficient, etc.) on the bootstrap sample
- Repeat  $B$  times, forming a distribution of the statistic estimates  $\{\hat{\theta}_1^*, \dots, \hat{\theta}_B^*\}$

**Bootstrap SE/CI:**

- Standard error  $\approx \text{SD}(\hat{\theta}_b^*) = \sqrt{\frac{1}{B-1} \sum_{b=1}^B (\hat{\theta}_b^* - \bar{\theta}^*)^2}$
- Use percentiles or normal approximation to construct confidence intervals
- Interpreting the coverage of confidence intervals requires care

**Key premise:** The observed sample is representative of the population

# Review: Subset selection

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**Goal:** Identify a relevant subset of predictors among many

**Best subset selection:**

- Tries all  $2^p$  subsets (exhaustive); picks the best model for each size  $k$ , then chooses among them by adjusted  $R^2$ , CV, etc.
- Feasible only if  $p$  is small or moderate (can be very expensive for large  $p$ )

**Forward/backward stepwise:**

- Greedy approximations: add/remove one predictor at a time
- Complexity  $\mathcal{O}(p^2)$  vs.  $2^p$  for best subset
- Might miss the absolute best subset but often works well in practice

**Pros/Cons:**

- Direct variable selection (some coefficients set to zero)
- Can be unstable for large  $p$ ; small changes in data may change chosen subset

# Review: Regularization

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**Motivation:** Least squares can be unstable or undefined if  $p \approx n$  or  $p > n$ ; high variance or collinearity issues

**Ridge regression:**

- Add penalty  $\lambda \sum_j \beta_j^2$
- Typically shrinks all coefficients; no exact zeros
- More stable under collinearity

**Lasso:**

- Add penalty  $\lambda \sum_j |\beta_j|$
- Can zero out some coefficients, enabling variable selection
- Slightly less stable than ridge if predictors are highly correlated

**Tuning  $\lambda$ :** Usually chosen by cross-validation; neither ridge nor lasso always wins—depends on data and interpretability needs

# Review: Multiple hypotheses testing

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**Problem:** Testing many hypotheses inflates chance of false positives

- Probability( $\geq 1$  false positive) can be  $1 - (1 - \alpha)^m$  if tests are independent
- $p$ -hacking: repeatedly searching for small  $p$ -values leads to spurious “discoveries”

**FWER** (Family-Wise Error Rate):

- Probability of any (=at least 1) false positive
- Bonferroni, Holm's step-down keep  $\text{FWER} \leq \alpha$
- Often conservative, can reduce power when  $m$  is large

**FDR** (False Discovery Rate):

- Expected fraction of false positives among rejections (=FP + TP)
- Benjamini–Hochberg procedure can control FDR
- Less conservative, typically yields more rejections, tolerating some false positives